Spatial and temporal boundaries to gene flow between Chaenocephalus aceratus populations at South Orkney and South Shetlands

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ABSTRACT: The black-fin icefish Chaenocephalus aceratus is among the most abundant fish species on the Antarctic continental shelves of the Scotia Arc, and Bouvet Island. We genotyped 11 microsatellite loci in C. aceratus population samples from South Orkney, southern South Shetlands, and Elephant Island (northern South Shetlands) collected in 2002 and 2006. This investigation further develops a previous study on the species reporting the presence of one panmictic population in southern South Shetlands and Elephant Island, with genetic differentiation between year classes. Our results reveal a more complex pattern of differentiation than shown previously, as genetic differences occur both at the temporal level at Elephant Island and at the geographic scale between southern South Shetland–Elephant Islands and South Orkney population samples. In particular, the magnitude of genetic differentiation at the temporal scale, the relatively high effective population size ($N_e$) and high gene flow indicate that genetic differentiation is not only driven by geographic distance. At present, our results should be taken into account when defining conservation measures and management boundaries in regions where fishery is still open or where other Antarctic fish species are still exploited.

KEY WORDS: Chaenocephalus aceratus · Icefish · Southern Ocean · Microsatellite · Automatic allele calling · Geographic and temporal differentiation · Gene flow · $N_e$

INTRODUCTION

The black-fin icefish Chaenocephalus aceratus (Perciformes, Notothenioidei, Channichthyidae) is distributed in the Scotia Sea (western Atlantic Sector of the Southern Ocean) along the Scotia Arc. The species is also present further east at Bouvet Island (eastern Atlantic Sector of the Southern Ocean) (Kock & Jones 2002). Larvae of C. aceratus were also found near Arthur Harbor (Anvers Island, Antarctic Peninsula [AP]; Hemmingsen & Douglas 1970) and, although this finding has never been further confirmed, this would extend the distribution further west and south. The Scotia Sea extends from 20° to 65° W and is bounded by the Weddell Sea to the south, the South Atlantic Ocean to the north and east, and the Pacific Ocean to the west of Drake Passage (our Fig. 1; Brandon et al. 2004). The regional circulation is dominated by the eastward flow of the Antarctic Circumpolar Current (ACC) (Orsi et al. 1995) that enters the Scotia Sea through Drake Passage (Cunningham et al. 2003). There is also a strong northward outflow of waters from the Weddell Sea occurring at a depth interval between 2000 and 4000 m (Naveira Garabato et al. 2002, Schodlok et al. 2002). As exemplified by drifters and modeling of krill transport (Fach & Klinck 2006), the ACC current creates a preferential passive dispersal route from southwest sites to the northeast.
ones, potentially creating a source-sink relationship between populations of organisms inhabiting the area.

The continental shelves of the AP and Scotia Arc Islands host a rich biological community (Moffat et al. 2008) where *Chaenocephalus aceratus* is among the most abundant fish species (Kock & Stransky 2000). Adults occur at depths of 200 to 400 m, although some individuals have been found beyond 700 m depth (Kock & Jones 2002, Kock 2005). From a commercial point of view, the black-fin icefish represented an important fishery resource until the 1989–1990 season (Kock 1992). After a deep exploitation of notothenioid stocks, commercial finfish harvest was banned in 1990–1991 by the Commission of the Convention for the Conservation of Antarctic Marine Living Resources (CCAMLR) along the Southern Scotia Arc and restricted to South Georgia and South Sandwich Islands (CCAMLR Conservation Measures 32-03 and 32-04).

Kock & Jones (2002) observed a remarkable fluctuation in frequencies of age classes of *Chaenocephalus aceratus* over shelf areas of South Shetlands, including Elephant Island, and across most sectors of the South Orkneys shelf (Kock & Stransky 2000, Kock & Jones 2005). In fact, while in some years, Age Class I (1 yr old individuals) was very frequent in bottom-trawl hauls, in other years the youngest catches included only individuals from Age Class II. Kock & Jones (2002) suggested that changes in feeding conditions between years causes differences in the mid-water distribution of these fish (and consequently its abundance in bottom trawls) from year to year. *C. aceratus* shows a developmental change in feeding habit, from larvae that feed on furcilia (larval krill stages), to juveniles that mainly feed on mysids and krill. After having become 3 to 4 yr old, mature adults change their diet and display a benthic sit-and-wait feeding strategy, mostly targeting fish (Kock et al. 2000). Only occasionally do adults ascend to near-bottom or mid-water layers to feed on krill (Siegel 1980, Takahashi & Iwami 1997).

While most studies on *Chaenocephalus aceratus* concern mainly growth, ecology, life history, or behavioral aspects (see e.g. Flores et al. 2004, La Mesa et al. 2004, Detrich et al. 2005), recently a few investigations have focused on the population genetics of the species (Papetti et al. 2007, Susana et al. 2007). In waters of the Southern Ocean, understanding the role of factors that may influence population genetic structure has gained
additional attention due to the increasing need of proper management of natural resources (FAO 2001). Failure to detect hidden stock units can lead to local over-fishing and in extreme cases to severe declines in species abundance (Frank & Brickman 2001). This also means that temporal monitoring of stock genetic structure should be implemented as a further strategy in management plans in areas still exploited, despite the difficult accessibility of Southern Ocean to obtain replicates from different years.

Our study follows a previous investigation on Chaenocephalus aceratus, where one panmictic population was found at Elephant Island and southern South Shetlands (Papetti et al. 2007). This result is in agreement with parasitic infestation patterns and data on spawning time/maturity stages that documented the presence of a single population on the microgeographic scale of South Shetlands (Kock & Möller 1977, Siegel 1980). Studies on parasite infestation, however, suggested the existence of at least 3 differentiated stocks off the shelves of South Georgia, South Orkney, and South Shetland Islands (Siegel 1980, Sosinski & Janusz 2000), prompting the analysis of new population samples. Genetic markers also evidenced temporal variability in the genetic pool of C. aceratus (Papetti et al. 2007) as there was found to be significant differentiation between year classes. While studies on marine organisms usually assume genetic homogeneity of broadly distributed species (Ward et al. 1994), this view is challenged by the growing body of literature that revealed genetic differentiation at geographic scales (Patarnello et al. 2007), and recently by several studies documenting genetic differentiation at the temporal scale, for instance in Gadus morhua, Merluccius merluccius, Pleuronectes platessa, and Clupea harengus (Ruzzante et al. 1999, Lundy et al. 2000, Hoarau et al. 2002, Nielsen et al. 2003, Bekkevold et al. 2005, Hede Jørgensen et al. 2005, O’Leary et al. 2006). For example, Lundy et al. (2000) reported temporal differentiation among hake (M. merluccius) population samples collected from the Bay of Biscay despite temporally stable spawning sites and juvenile retention. In particular, this suggests that temporal changes at microsatellite allele frequency were far more important than geographic distances between populations (Lundy et al. 2000), similarly to what is observed in the European eel (Dannewitz et al. 2005, Maes et al. 2006).

Large variance in reproductive success and temporal spawning patterns, low effective population size (N_e), fluctuations in recruitment, selection on early life stages, and gene flow from genetically differentiated populations are common in marine fish and may all act in preventing genetic homogeneity between year classes or temporal replicates, and in maintaining unstable population units (McQuinn 1997). The ‘fluctuating genetic patchiness’ defines the mix of these factors that contribute to genetic differentiation among temporal replicates from the same location (Hauser et al. 2002, Hellberg et al. 2002, Hedgecock et al. 2004, Hede Jørgensen et al. 2005).

In this context, the present study is aimed at verifying whether: (1) the genetic structure of Chaenocephalus aceratus is stable over time in previously analyzed sampling areas, (2) there are differentiated population units on a larger geographic scale, and (3) levels of differentiation are divergent at a geographic/temporal scale. In order to address these points, we genotyped 11 microsatellite loci in 4 C. aceratus population samples from 3 locations (South Orkney, southern South Shetland, and Elephant Islands) collected during different years (2002 and 2006). This is, to our knowledge, the first attempt to dissect temporal and geographic variation in a notothenioid species using microsatellites. While our study concerns samples collected in areas closed to finfish harvest (CCAMLR Conservation Measures 32-03 and 32-04), we expect that our results can be useful for future management plans in still-exploited areas of the Scotia Sea such as South Georgia and South Sandwich Islands.

MATERIALS AND METHODS

Sampling. Population samples analyzed in the present study were collected at 3 different sites of the Scotia Sea, namely Elephant Island, southern South Shetlands, and South Orkneys (Table 1, Fig. 1a); at Elephant Island, samples were obtained from similar locations in 2002 and 2006. In the present study, population samples and temporal replicates from South Shetland Islands (King George, Livingston, and Elephant Islands) are conventionally defined as ‘southern South Shetlands and Elephant Island’ populations, due to their coordinates origin (our Table 1; see also Papetti et al. 2007 and their online supplementary materials). A small piece of muscle tissue was collected from each specimen and preserved in absolute ethanol until the molecular analysis. For the 2006 population samples, age information was not available; for this reason analysis considering age classes differentiation was not performed.

Genetic and statistical analyses. DNA extraction and genotyping: Total genomic DNA was extracted from 10 to 100 mg of absolute ethanol-preserved muscle tissue from 95 specimens collected at South Orkney and Elephant Islands (Table 1) according to Patwary et al. (1994). Eleven microsatellite loci were amplified as described in Papetti et al. (2007). Seven of these were species-specific (Susana et al. 2007) while 4 loci,
isolated from *Chionodraco rastrospinosus* (Notothenioidei, Channichthyidae) (Papetti et al. 2006), cross-amplified in *Chaenocephalus aceratus*. Fragment analysis was performed at the BMR Genomics Molecular Biology service (www.bmr-genomics.it). In order to perform a comparable allele sizing between our previous study and the present study, 10 individuals collected during 2002 and already genotyped in Papetti et al. (2007) were randomly chosen and re-genotyped so as to test the consistency of allele amplification and sizing. Moreover, in order to negate the negative consequences of a poor allele calling, binning was automated with the software **FLEXIBIN** ver. 2 (Amos et al. 2007) and all samples from Papetti et al. (2007) were rescoring. Allele calling is known to be prone to a high error rate when new alleles are found, where reference standards are not available or, as in the present case, when scoring new samples (Amos et al. 2007). The scoring was then manually checked by the authors and loci were analyzed for null alleles presence with **MICROCHECKER** ver. 2.2.3 (van Oosterhout et al. 2004). Our procedure allowing checking for discrepancies between present and previous scoring and results from the previously manually scored alleles were confirmed using automated scoring.

**Genetic diversity, Hardy-Weinberg equilibrium, linkage disequilibrium:** Descriptive statistics such as allele size range (\(S_a\)) in base pair (bp), number of alleles (\(N_a\)), and most common allele frequency at each locus were computed with **POWERMARKER** ver. 3.25 (Liu & Muse 2005). Observed heterozygosity (\(H_O\)) and unbiased gene diversity (\(H_B\)) were calculated using the Windows Excel add-in package **GENALEx** ver. 6 (Peakall & Smouse 2006). Genotypic linkage disequilibrium tests between pairs of loci in each population and global tests for conformity with Hardy-Weinberg equilibrium were performed across loci and across populations using **GENEP OP**, online version (Raymond & Rousset 1995).

**Population structure:** Differentiation tests between population samples were performed using **CHIFISH** ver. 1.3 (Ryman 2006). The software tests the hypothesis of no allele frequency difference among populations at each locus by means of Pearson’s traditional chi-squared and Fisher’s exact test. **FSTAT** ver. 2.9.3.2 (Goudet 1995) was used to compute the allelic richness (\(A_R\)) and the overall and population pairwise \(F_{ST}\). The \(A_R\) was based on the smallest sample size between populations, of 45 diploid individuals (population from South Shetlands). The estimation of the probability of \(F_{ST}\) was determined using 1000 permutations and bootstrap replicates for all comparisons. A standardized measure of population divergence (\(F^*_{ST}\)), independent from the levels of heterozygosity, was calculated to allow comparisons between populations (Hedrick 2005). The program **RECODEDATA** ver. 1.0 (available from www.bentleydrummer.nl/software/index.html) was used to recode the data set so that each population sample had unique alleles, allowing the estimation of the maximum level of population differentiation. To estimate the standardized level of population subdivision (\(F^*_{ST}\)), each actual estimate of population differentiation (\(F_{ST}\)) was divided by the maximum \(F_{ST}\) value obtained from the recoded data set. The 95% confidence intervals were estimated by 1000 bootstrap replicates over loci and probability values were determined by 1000 permutations. For all our tests with multiple comparisons, statistical significance level was adjusted, against type I errors, using standard Bonferroni correction (Rice 1989). The nominal significance level was set at 0.05.

Population structure at the geographic level was investigated by a landscape genetic approach using the program **BARRIER** ver. 2.2 (Manni et al. 2004) to find the largest breaks in genetic structure of our populations collected at different locations (South Orkney, South Shetland, and Elephant Islands). Using a Delaunay triangulation obtained with the Voronoi tessella-

### Table 1. *Chaenocephalus aceratus*. Population samples used in this study. N: sample size

<table>
<thead>
<tr>
<th>Population sample geographic origin</th>
<th>Acronym</th>
<th>Sampling year</th>
<th>Sampling coordinates</th>
<th>Sampling cruise</th>
<th>N</th>
</tr>
</thead>
<tbody>
<tr>
<td>Elephant Island</td>
<td>EI02</td>
<td>2002</td>
<td>–</td>
<td>ANT-XIX/3 AWI(^b)</td>
<td>202</td>
</tr>
<tr>
<td></td>
<td>EI06</td>
<td>2006</td>
<td>61° 04’ 35&quot; S, 56.01° 10’ W</td>
<td>JR145 BAS(^c)</td>
<td>47</td>
</tr>
<tr>
<td>Southern South Shetlands</td>
<td>SS02</td>
<td>2002</td>
<td>–</td>
<td>ANT-XIX/3 AWI(^b)</td>
<td>45</td>
</tr>
<tr>
<td>South Orkney</td>
<td>SO06</td>
<td>2006</td>
<td>60° 54’ 22&quot; S, 46° 30’ 59” W</td>
<td>JR145 BAS(^c)</td>
<td>48</td>
</tr>
</tbody>
</table>

\(^a\)See Table 1S in Papetti et al. (2007) online supplementary materials

\(^b\)CCAMLR fish survey and ANDEEP I cruise PFS ‘Polarstern’ ANT-XIX/3, AWI: Alfred Wegener Institute, Bremerhaven, Germany; population sample previously analyzed in Papetti et al. (2007)

\(^c\)JR145 cruise with RRS ‘James Clark Ross’, BAS, British Antarctic Survey, samples collected under NERC AFI 6/16 grant to G. Carvalho and kindly provided by J. Rock and T. North
tion estimator (Manni et al. 2004), BARRIER projects a geometrical representation of the populations. As output, the software provides the putative genetic barriers starting with the pair of populations with the largest genetic distance using the Monmonier maximum differential algorithm (Monmonier 1973). Robustness of the barriers is determined by re-sampling the distance matrices 100 times. Two datasets were analyzed, excluding each time 1 of the 2 temporal replicates (Elephant Island 2002 and 2006). As a parallel approach to graphical representation of genetic barriers, we tested for the presence of correlation between geographic and genetic distance performing a Mantel’s test using the procedure of Smouse (1986) implemented within the Windows Excel add-in package GenAlEx ver. 6 (Peakall & Smouse 2006). Statistical significance of the values was obtained via 9999 random permutations.

**Gene flow and effective population size:** The effective number of migrants per generation \(N_{e,m}\), population effective size per migration rate \(m\) was estimated for the Elephant Island 2006–South Orkneys 2006 (E106/ SO06; for complete list of acronyms see Table 1) pair of populations because they represent 2 different locations sampled at the same time (2006). \(N_{e,m}\) for this populations pair was obtained using 2 methods: (1) Wright’s formula: \(F_{ST} = 1/(1 + 4N_{e,m})\) (Wright 1965) and (2) Slatkin’s private allele method (Slatkin 1985) with the correction for sample size (Barton & Slatkin 1986) as implemented in GenePop, online version (Raymond & Roussset 1995).

Although the standard approach to measure gene flow (and migration rates) is through the use of \(F_{ST}\) (Wright 1965, Neigel 2002), recent developments have produced population genetic software programs with more sophisticated approaches that take into account more of the underlying biology of populations and thereby produce more accurate information (Abdo et al. 2004, Pearse & Crandall 2004). In fact, although \(F_{ST}\) is theoretically related to equilibrium migration rates, real populations are very likely to violate the assumptions of the model. Similar \(F_{ST}\) values, for example, might reflect both low gene flow and large \(N_{e}\), or high gene flow and low \(N_{e}\) (Whitlock & McCauley 1999, Jensen et al. 2005, Fraser et al. 2007). For these reasons, we calculated migration rates (as \(N_{e,m}\)) by estimating \(M = m/\mu\) (migrants per generations scaled by mutation rate) and \(\theta = 4N_{e}\mu\) values with MIGRATE ver. 2.1.3 software (Beerli & Felsenstein 1999, 2001). MIGRATE provides a long-term method of \(\theta\) and \(M\) estimation based on coalescent theory (Beerli & Felsenstein 2001) under the assumptions of migration-drift equilibrium, populations constant sizes and gene flow over the coalescent period (\(\approx 4N_{e}\) generations) (Fraser et al. 2007). MIGRATE was successively run using as starting values the estimates obtained from previous runs and the following parameters: microsatellites Ladder model (Ohta & Kimura 1973, Kimura & Ohta 1978), 10 short chains (10 000 sampled, 500 recorded) and 2 long chains (4 000 000 sampled, 200 000 recorded; number of discarded trees per chain: 400 000). The software was run 2 times to assure that final chains were estimating the same values of \(\theta\) and \(M\). The results obtained from different runs were very similar; thus we report only 1 set of results. Values of \(\theta\) were converted to \(N_{e}\) using mutation rates (\(\mu\)) of \(5 \times 10^{-4}\) (Estoup & Angers 1998).

Effective population size \(N_{e}\) was also calculated with a heterozygosity-based method (Ohta & Kimura 1973), which predicts that at mutation-drift equilibrium under the stepwise mutation model (SMM), \(N_{e}\) should equal \([1/(1–H)^2–1]/8\mu\), where \(H_{e}\) is the observed averaged heterozygosity and the mutation rate, \(\mu\), was assumed to be \(5 \times 10^{-4}\) according to Estoup & Angers (1998). Further estimation of \(N_{e}\) was performed using NEEstimator ver. 1.3 (Peel et al. 2004). The software provides point estimation of \(N_{e}\) for single populations; in the present study, the linkage disequilibrium (LD) method by Bartley et al. (1992) was chosen. In particular, the assumptions necessary to estimate \(N_{e}\) from correlations of alleles include neutral alleles, no migration, no subpopulation structure, and a random sampling of the entire population (Hill 1981). The software calculates LD and the correlation among alleles at different loci \(r\) in a population sample and takes the arithmetic mean of all \(r^2\) values to get a single \(r^2\) which is used to calculate \(N_{e}\) as in the equation \(N_{e} = 1/[3(r^2 × 1/S)]\) (\(S = \) sample size) (Campton 1987, Bartley et al. 1992, Peel et al. 2004, Fraser et al. 2007).

**RESULTS**

**Genetic diversity, Hardy-Weinberg equilibrium, linkage disequilibrium**

The number of alleles per locus ranged from 6 (Ca88) to 93 (Ca21; Table 2). Frequencies of the most common allele (MCAF) at each locus were similar between population samples, except for locus Cr259 that showed the frequency of 0.183 in the South Orkneys population sample compared to the frequencies observed in the other populations that ranged between 0.355 and 0.383. Considering all samples together, the observed heterozygosity ranged from 0.157 to 0.709, while the expected heterozygosity ranged from 0.208 to 0.945. A weak excess of homozygotes was found at all loci, except for Ca48 (Table 2).

Hardy-Weinberg equilibrium (HWE) probabilities (Table 2) were calculated for each locus, for each population and combined by Fisher’s test for a global value.
Table 2. Chaenocephalus aceratus. Summary statistics for genotyping: allelic variability at 11 microsatellite loci in 4 population samples and in all populations. Reported are: allele size range (S), number of alleles (N), most common allele frequency, allelic richness (R), observed heterozygosity (O)/unbiased gene diversity (E) at each locus. In the last 5 columns: probabilities of Hardy-Weinberg equilibrium (pHWE) following the sequential Bonferroni correction (nominal significance threshold \( \alpha = 0.05 \), comparisons number \( k = 4 \); significant p-values in **bold**). TOT: total; mean: average values across all loci for all summary statistics.

<table>
<thead>
<tr>
<th>Locus</th>
<th>S</th>
<th>Population samples</th>
<th>TOT Population samples</th>
<th>Average allele frequency</th>
<th>H0/H0</th>
<th>pHWE</th>
</tr>
</thead>
<tbody>
<tr>
<td>Ca21</td>
<td>202– 304</td>
<td>27 42 37 93</td>
<td>0.162 0.166 0.177 0.132 0.1799</td>
<td>2.908 2</td>
<td>2.894 2</td>
<td>2.61 0.840/ 0.755/ 0.819/ 0.770/ 0.651/ 0.0025 0.0009 0.0021 0.0007 p &lt; 0.0001</td>
</tr>
<tr>
<td>Ca26</td>
<td>153– 284</td>
<td>25 19 16 14 28</td>
<td>0.296 0.300 0.302 0.369 0.28506</td>
<td>31.342 26.51</td>
<td>39.838 34.48</td>
<td>39.87 0.178/ 0.200/ 0.222/ 0.250/ 0.170/ 0.208 0.3191</td>
</tr>
<tr>
<td>Ca40</td>
<td>154– 216</td>
<td>44 30 28 32 49</td>
<td>0.138 0.088 0.081 0.141 0.12126</td>
<td>16.596 18.459</td>
<td>15.77 13.659</td>
<td>16.624 0.184/ 0.184/ 0.184/ 0.184/ 0.184/ 0.0025 0.0009 0.0021 0.0007 p &lt; 0.0001</td>
</tr>
<tr>
<td>Ca48</td>
<td>160– 216</td>
<td>5 5 5 7</td>
<td>0.903 0.900 0.863 0.877 0.89412</td>
<td>29.048</td>
<td>26.59</td>
<td>27.542 39.775 0.178/ 0.200/ 0.222/ 0.250/ 0.170/ 0.208 0.2196 1.0000 0.0842 1.0000 0.4943</td>
</tr>
<tr>
<td>Ca55</td>
<td>263– 329</td>
<td>26 21 22 19 28</td>
<td>0.127 0.122 0.119 0.160 0.12647</td>
<td>3.939</td>
<td>4.863</td>
<td>4.929</td>
</tr>
<tr>
<td>Ca66</td>
<td>73– 181</td>
<td>42 22 19 22 44</td>
<td>0.138 0.159 0.200 0.177 0.14418</td>
<td>20.831</td>
<td>20.655</td>
<td>21.242</td>
</tr>
<tr>
<td>Ca88</td>
<td>109– 148</td>
<td>15 14 16 16 6</td>
<td>0.167 0.197 0.163 0.153 0.18828</td>
<td>25.387</td>
<td>21.677</td>
<td>18.551</td>
</tr>
<tr>
<td>Cr15</td>
<td>142– 198</td>
<td>5 2 3 2 19</td>
<td>0.035 0.031 0.073 0.087 0.04097</td>
<td>14.074</td>
<td>13.907</td>
<td>15.677</td>
</tr>
<tr>
<td>Cr127</td>
<td>104– 126</td>
<td>9 7 7 7 9</td>
<td>0.425 0.477 0.336 0.469 0.42696</td>
<td>7.215</td>
<td>6.933</td>
<td>6.894</td>
</tr>
<tr>
<td>Cr171</td>
<td>171– 260</td>
<td>40 26 30 38 59</td>
<td>0.114 0.154 0.119 0.082 0.089677</td>
<td>24.642</td>
<td>26</td>
<td>28.966</td>
</tr>
<tr>
<td>Ca269</td>
<td>211– 300</td>
<td>21 16 17 19 22</td>
<td>0.355 0.383 0.380 0.183 0.33728</td>
<td>17.632</td>
<td>15.86</td>
<td>16.445</td>
</tr>
</tbody>
</table>

Mean 26.18 17.18 18.64 19.18 33.09 0.333 0.352 0.321 0.333 0.32359 17.59 16.93 18.13 18.39 18.69 0.715/ 0.704/ 0.716/ 0.688/ 0.564/ 0.768 0.746 0.777 0.758 0.762 |
Elephant Island (2002) and South Orkneys (2006) ($F_{ST}$ = 0.011) and southern South Shetlands (2002) and South Orkneys (2006) ($F_{ST}$ = 0.015), whereas the level of genetic differentiation between temporal replicates appears slightly lower ($F_{ST}$ = 0.009 between Elephant Island 2002 and 2006). This is also true for standardized $F_{ST}$ values (Table 4). A significant relationship between genetic and geographic distance matrices was confirmed by Mantel's test, alternatively considering 1 of the 2 temporal replicates from Elephant Island (dataset with EI02, $R^2$ = 0.1319; $p < 0.00001$, after 9999 permutations; dataset with EI06, $R^2$ = 0.1441; $p < 0.00001$, after 9999 permutations). Barriers to the gene flow were graphically represented by BARRIER ver 2.2

(Manni et al. 2004) (Fig. 2), which pointed out a large genetic break (namely Barrier I in Fig. 2) between South Orkney and the southern South Shetlands/Elephant Island.

### DISCUSSION

The present study reveals a much more complex pattern of differentiation than shown previously in Papetti et al. (2007), as genetic differences occur both at temporal (Elephant Island 2002 and 2006) and at geographic scale between South Shetlands (including Elephant Island) and South Orkney. Interestingly, similar conclusions on geographic differentiation were also reached by parasitic infestation patterns and data on spawning time/maturity stages (Siegel 1980, Sosinski & Janusz 2000).

The present study, together with the rising number of recent cases such as Gadus morhua and Platichthys flesus (Pogson et al. 2001, Hemmer-Hansen et al. 2007), demonstrates that potentially high effective population size and absence of obvious barriers to

### Table 3. Chaenocephalus aceratus. Differentiation probability test between population samples. Multilocus p-values were estimated by Fisher’s and chi-squared methods ($p(\text{Fisher})/p(\chi^2)$). *Significant p-values after sequential Bonferroni adjustment. n loci: number of loci showing significant differences for each differentiation test; all: all loci; without: excluding loci that deviated from Hardy-Weinberg equilibrium

<table>
<thead>
<tr>
<th>Population sample pairs</th>
<th>Loci</th>
<th>p (Fisher)</th>
<th>n loci</th>
<th>$\chi^2$</th>
<th>df</th>
<th>$p (\chi^2)$</th>
<th>n loci</th>
</tr>
</thead>
<tbody>
<tr>
<td>All populations</td>
<td>All</td>
<td>&lt;0.00001*</td>
<td>7</td>
<td>1965.232</td>
<td>1059</td>
<td>&lt;0.00001</td>
<td>6</td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>0.00007*</td>
<td>3</td>
<td>539.707</td>
<td>402</td>
<td>0.00001</td>
<td>2</td>
</tr>
<tr>
<td>EI02/SS02</td>
<td>All</td>
<td>0.10040</td>
<td>1</td>
<td>328.64</td>
<td>291</td>
<td>0.06363</td>
<td></td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>0.23417</td>
<td>0</td>
<td>137.493</td>
<td>125</td>
<td>0.20074</td>
<td></td>
</tr>
<tr>
<td>EI02/EI06</td>
<td>All</td>
<td>&lt;0.00001*</td>
<td>3</td>
<td>754.518</td>
<td>314</td>
<td>&lt;0.00001</td>
<td>5</td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>0.01968</td>
<td>1</td>
<td>163.731</td>
<td>126</td>
<td>0.01339</td>
<td>2</td>
</tr>
<tr>
<td>EI02/SO06</td>
<td>All</td>
<td>&lt;0.00001*</td>
<td>7</td>
<td>797.415</td>
<td>322</td>
<td>&lt;0.00001</td>
<td>7</td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>0.00050*</td>
<td>3</td>
<td>187.361</td>
<td>125</td>
<td>0.00026</td>
<td>3</td>
</tr>
<tr>
<td>EI06/SS02</td>
<td>All</td>
<td>&lt;0.00001*</td>
<td>6</td>
<td>430.929</td>
<td>254</td>
<td>&lt;0.00001</td>
<td>5</td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>0.00007*</td>
<td>4</td>
<td>150.458</td>
<td>100</td>
<td>0.00083</td>
<td>4</td>
</tr>
<tr>
<td>EI06/SO06</td>
<td>All</td>
<td>&lt;0.00001*</td>
<td>6</td>
<td>385.98</td>
<td>261</td>
<td>&lt;0.00001</td>
<td>5</td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>0.00131*</td>
<td>2</td>
<td>140.215</td>
<td>97</td>
<td>0.00272</td>
<td>2</td>
</tr>
<tr>
<td>EI02+EI06/SO06</td>
<td>All</td>
<td>&lt;0.00001*</td>
<td>7</td>
<td>398.246</td>
<td>219</td>
<td>&lt;0.00001</td>
<td>6</td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>&lt;0.00001*</td>
<td>5</td>
<td>229.457</td>
<td>82</td>
<td>&lt;0.00001</td>
<td>5</td>
</tr>
<tr>
<td>EI02+EI06/SS02</td>
<td>All</td>
<td>&lt;0.00001*</td>
<td>2</td>
<td>1212.954</td>
<td>652</td>
<td>&lt;0.00001</td>
<td>3</td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>0.11040</td>
<td>0</td>
<td>153.299</td>
<td>130</td>
<td>0.0796</td>
<td>0</td>
</tr>
<tr>
<td>SS02/SS06</td>
<td>All</td>
<td>&lt;0.00001*</td>
<td>5</td>
<td>431.518</td>
<td>254</td>
<td>&lt;0.00001</td>
<td>4</td>
</tr>
<tr>
<td></td>
<td>Without</td>
<td>0.00106*</td>
<td>2</td>
<td>139.331</td>
<td>97</td>
<td>0.00316</td>
<td>2</td>
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</tbody>
</table>

### Table 4. Chaenocephalus aceratus. Estimation of the genetic differentiation based on allele frequencies. $F_{ST}$: actual estimate of population differentiation and 95% confidence interval ($F_{ST}$ 95% CI); $F_{ST}$*: standardized measure of population divergence

<table>
<thead>
<tr>
<th>Sample pair</th>
<th>$F_{ST}$</th>
<th>95% CI</th>
<th>$F_{ST}$*</th>
</tr>
</thead>
<tbody>
<tr>
<td>EI02/SS02</td>
<td>0.002</td>
<td>-0.001 to 0.006</td>
<td>0.0085</td>
</tr>
<tr>
<td>EI02/EI06</td>
<td>0.009</td>
<td>0.001 to 0.020</td>
<td>0.0404</td>
</tr>
<tr>
<td>EI02/SO06</td>
<td>0.011</td>
<td>0.004 to 0.020</td>
<td>0.0478</td>
</tr>
<tr>
<td>EI06/SS02</td>
<td>0.014</td>
<td>0.005 to 0.025</td>
<td>0.0611</td>
</tr>
<tr>
<td>EI06/SO06</td>
<td>0.009</td>
<td>0.005 to 0.013</td>
<td>0.0403</td>
</tr>
<tr>
<td>EI02+EI06/SO06</td>
<td>0.008</td>
<td>0.004 to 0.014</td>
<td>0.0358</td>
</tr>
<tr>
<td>EI02+EI06/SS02</td>
<td>0.003</td>
<td>0.000 to 0.007</td>
<td>0.0131</td>
</tr>
<tr>
<td>SS02/SS06</td>
<td>0.015</td>
<td>0.005 to 0.024</td>
<td>0.0630</td>
</tr>
</tbody>
</table>

Levels of gene flow between populations are reported in Table 5 as $N_{m}$ and were calculated following different approaches. Results range from 8.20 to 49.75 migrants per generation indicating a high general gene flow trend between the 2 geographic areas. In particular, while Wright’s and Slatkin’s methods provide unidirectional gene flow values only, MIGRATE ver. 2.1.3 allows calculation of migration into 2 separated gene-flow directions and indicated a rather symmetric gene flow between Elephant Island and South Orkneys in 2006 (Fig. 1b, Table 5). In agreement with previously published studies, effective population size ($N_{e}$) varied considerably between methods (Fraser et al. 2007), with MIGRATE providing the highest estimates (Table 6); the $N_{e}$ values are nevertheless different from previous estimates calculated between cohorts considering drift only ($N_{e}$ = 96 in Papetti et al. 2007). In particular, Elephant Island and South Orkney showed higher $N_{e}$ values than before, also when considering the 95% confidence intervals (average from 241.4 to 21,236.39 for EI06 and from 264.00 and 16,491.42 for SO06; Table 6).
gene flow (Ward et al. 1994) cannot prevent population structuring in marine fishes on geographical scales. In several marine species such as G. morhua, Sebastes rastrelliger, and Caribbean reef fish, the extent of gene flow predicted from larval development, motility, or time spent as planktonic stage does not, in fact, always match with the observed population structure (Hedgecock 1986, Pogson et al. 2001, Taylor & Hellberg 2003, Buonaccorsi et al. 2004).

In the present study, the Migrate coalescent approach to estimate directional migration rates has been applied for the first time to an Antarctic fish, suggesting gene flow against the dominant currents from South Orkney to Elephant Island (15.51°N em). This result can be explained by adults' backward migration against the ACC and the northward outflow of waters from the Weddell Sea. If confirmed, this hypothesis would imply, from a management perspective, that populations from still-exploited areas might contribute to the genetic pool of areas where fishing is banned. However, considering that the southern South Shetlands/ Elephant Island and South Orkney areas of the Antarctic shelf are separated by a deep-water channel (Philip Passage, ~2000 m depth; Schodlok et al. 2002, Brandon et al. 2004) and taking into account the sedentary habitat of Chaenocephalus aceratus adults (Jones et al. 2001), this result needs further investigations. In particular, it is worth noting that unsampled southwesterly or southeasterly populations might contribute immigrants to both Elephant Island and South Orkney populations, producing the observed gene flow pattern. Thus, a more thorough sampling in the distribution area of C. aceratus will be crucial to defining individual origin and gene-flow patterns.

On the other hand, the present study revealed a complex differentiation pattern in which temporal genetic differences also play an important role. Together with some examples from both...
unexploited and commercially exploited species (Ruzante et al. 1999, 2006, Nielsen et al. 2003, Bekkevold et al. 2005, Hede Jørgensen et al. 2005, O’Leary et al. 2006), the present study revealed that significant temporal genetic heterogeneity may occur in a marine fish species. In particular, for *Chaenocephalus aceratus*, the magnitude of genetic differentiation at a temporal scale (between temporal replicates collected at Elephant Island, $F_{ST} = 0.009$, $p < 0.0001$) indicates that genetic differentiation is not only driven by geographic distance. This result confirms our previous finding of a significant genetic differentiation between year classes (Papetti et al. 2007), indicating the existence of population structure over time. In that case, it was postulated that either recruitment and reproductive success variability, small effective population size, or migration may cause differences over time, while other forces, such as selection, can drive genetic differentiation.

**Recruitment and reproductive success variability**

Hedgecock (1994) proposed that temporal genetic variation in marine species may result from a large variance in reproductive success among adults driven by the unpredictability of oceanographic conditions (‘sweepstakes’ hypothesis, resulting in major discrepancies between observed and effective population sizes). This implies that the resulting cohorts comprise the progeny of a small proportion of the spawning population and may result in significant differences in the genetic composition of recruits over time.

Asynchronies between reproductive activity and suitable environmental conditions could lead to individual reproductive failure as a result of sperm limitation, variation in the availability of food for larvae (match-mismatch hypothesis; Cushing 1975), unpredictable nearshore oceanographic features (membranocavitation hypothesis; Sinclair 1988), and predation. The reproductive failure by a significant fraction of the adult population can result in temporal genetic variation (Maes et al. 2006) and a Wahlund effect, which occurs when genetically differentiated populations mix, leading to a reduction of observed heterozygosity (Wahlund 1928, Hoarau et al. 2002, 2007, but see Nielsen et al. 2003). In fact, a general deficit of heterozygotes found in the present study supports this scenario in *Chaenocephalus aceratus*. In addition, high recruitment variability was reported for this species (Kompowski 1990, Reid et al. 2007) and other Antarctic fish (Hill et al. 2005, Collins et al. 2007), suggesting the importance of variance in reproductive success in these organisms. In particular, for the mackerel icefish *Champsocephalus gunnari* in South Georgia (Hill et al. 2005) and the Patagonian toothfish *Dissostichus eleginoides* on the Shag Rocks shelf (Collins et al. 2007), the variability was explained by sea surface temperature (SST) fluctuations. Since *C. aceratus* and mackerel icefish are both channichthyids covering a similar latitudinal range and exhibiting some similar life history strategies, SST may play an important role in the determination of year-class strength also in *C. aceratus* (Reid et al. 2007) and, possibly, on its temporal genetic variation.

Differences in spawning time and place have also been implicated in temporal genetic differentiation (McPherson et al. 2003). Similarly to other notothenioids for which the onset of spawning exhibits latitudinal variation, such as *Dissostichus eleginoides* (Laptikhovsky et al. 2006) and the channichthid *Champsocephalus gunnari* (Eversen et al. 2001), *C. aceratus* populations inhabiting the southerly distribution range (South Shetlands) spawn from February to May, whereas northerly (South Georgia) spawning starts 1 or 2 mo earlier (Kock & Kellermann 1991). Moreover, mature individuals of this species are believed to move seasonally inshore to spawn in shallow coastal waters (<200 m), where they deposit eggs on the seabed and actively guard them (Detrich et al. 2005) until 4 mo, when larvae hatch (Kock & Jones 2002). Whereas the latitudinal breadth of the South Shetlands/South Georgia area is extreme and not comparable with the sampling area considered in the present study, shifts in reproduction times together with temporal variation in oceanographic conditions creates significant potential for temporal variation in recruitment and year-class distribution. In particular, following the predicted source-sink relationship based on regional circulation patterns (Fach & Klinck 2006), the temporal genetic differentiation observed at Elephant Island could be the result of a variable recruitment from populations located in the far southwest, in the Antarctic Peninsula.

Moreover, we cannot exclude the possibility that within the same spawning season, more than one spawning group of *Chaenocephalus aceratus* may occur in the same spawning ground at different times in ‘spawning waves’ as documented for instance in the Atlantic and Pacific herring (Lambert 1987). In these species, age structure determines the number of spawning waves per season in a population, with older herring arriving earlier in the season than smaller, younger ones (Lambert 1987, Hede Jørgensen et al. 2005). Different spawning events involving heterogeneous and possibly small groups of spawning individuals may therefore originate different assemblages of larvae by drift (So et al. 2006). This might represent an important reproductive strategy to maximize the survival rate of early life stages by spreading the effort.
over time to take advantage of (or to cope with) a variable environment (Lambert 1987), and it may partly explain the relationship between effective population size and observed census population size, as proposed by the ‘sweepstakes’ hypothesis (Hedgecock 1994).

Effective population size and migration

The present study provides the first estimations of long-term effective population size for Chaenocephalus aceratus. The combination of several analytical approaches, including MIGRATE and NEESTIMATOR, indicates average values of $N_e$ ranging from 241.4 to 21 236.39, with strong differences between methods (Table 6). These are long-term equilibrium estimates, calculated under simplified population models, and may become inaccurate if models’ assumptions are violated (Whitlock & McCauley 1999, Fraser et al. 2007). Taking into account these limitations, our $N_e$ values seem to indicate a relatively large population size in C. aceratus, which has been described as one of the most abundant species occurring in the Atlantic Ocean sector of the Southern Ocean (Kock & Stransky 2000). Considering that an effective population size of at least 50 individuals seems to be sufficient to prevent short-term loss of heterozygosity and 500 to maintain long-term adaptability (Frankham et al. 2002); our $N_e$ estimates seem to indicate that the species is not prone to genetic erosion at the evolutionary time-scale.

On the other hand, contemporary $N_e$ has been shown to be generally lower than its long-term counterpart (Fraser et al. 2007) and has been provisionally estimated, taking into account only differentiation between age classes, at about 100 individuals for Elephant Island in 2002 (Papetti et al. 2007). This discrepancy suggests the importance of non-equilibrium processes in determining temporal fluctuations, and prompts for a rigorous estimation of contemporary $N_e$ with temporal methods, which was not possible in the present study because of the presence of overlapping generations and the short time interval separating our temporal replicates (Jorde & Rynan 1995).

Similarly, it is interesting to consider our equilibrium estimates of gene flow (Fig. 1b, Table 5). Assuming an island model (Wright 1978), Waples & Gaggiotti (2006) suggested that a minimum of 5 or alternatively 25 effective migrants per generation are sufficient to prevent population differentiation (evolutionary population definition EV3 and EV4; Waples & Gaggiotti 2006). Thus, our gene-flow values ranging from 8.20 to 48.75 between different geographic locations should result in a single homogeneous population in the long run, and immigration from differentiated populations can hardly be invoked to explain temporal fluctuations at the same locality using this simplified gene-flow model.

While Chaenocephalus aceratus was expected to show a comparable differentiation pattern to that documented for many marine species with comparable high gene flow and $N_e$ (Hellberg et al. 2002), the present study indicated that C. aceratus is complexly structured, and a foremost subdivision boundary (temporal or geographic) between populations remains unclear. It is likely that shifts in spawning time and place, together with a certain degree of migration and a temporally variable complex stepping-stone scenario, generate a subtle population subdivision that is difficult to model with standard approaches.

In conclusion, our results suggest that Chaenocephalus aceratus populations may be genetically isolated from one another and, as fishery stocks, would deserve appropriate conservation policies. We also demonstrated a high level of genetic flow possibly acting against prevailing oceanographic conditions, across the off-shelf/deep-water habitat. Alternatively, this pattern can be produced by source populations contributing to both south Shetlands and South Orkney. In this sense, we cannot exclude the presence of additional differentiated populations in the AP that could also be responsible for the observed temporal variation. With the obvious difficulties in collecting samples in Antarctica, coordinated international efforts should be made to properly address these hypotheses.

At present, our results of geographic differentiation, high gene flow, and temporal variability should be taken into account when defining conservation measures and management boundaries in regions where the fishery is still open or where other Antarctic fish species are still exploited.

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