## SUPPLEMENT 1

## Gut contents analysis of English sole

## Collection Methods

In a concurrent but separate study by our research group, we examined the gut contents of English sole, and determined that there was no proportional change over time. We used 68 UWFC English sole specimens collected from 1930-2015 (SL = 141.553 $\pm 38.257 \mathrm{~mm}$ ). Twenty-three of these specimens were also used for our CSSIA-N analyses. In 2018, a UW research trawl collected an additional 76 English sole, and these specimens were preserved following UWFC protocol ( $\mathrm{SL}=155.066 \pm 40.306 \mathrm{~mm}$ ). Of these 76 , nine were randomly sampled to increase replication in the decade 2010-2020 (see Statistical analysis below) in the gut content analysis. Across UWFC stomachs and UW research trawl stomachs, we selected stomachs to sample by choosing those that appeared to be full. The contents were classified into 19 categories (amphipods, brachiopods, brittlestars, clams, copepods, corophid amphipods, crabs, Crangon spp., cumacea, isopods, mysids, polychaetes, scallops, snails, tubeworms, worms, unidentified digested material, empty). We recorded the mass of the categorized food items and total mass of the food items per fish in grams. None of the UW research trawl fish were used for CSSIAA-N. See Table S2 for UWFC accession numbers of fishes.

## Statistical Analysis

For each food item detected from our gut contents analysis, we conducted a beta regression. We selected this approach because response variables were a proportion of the total stomach content. Using this dataset ( $\mathrm{n}=75,1930-2018$ ) we examined proportional change in each prey item over time. To account for differences in fish size, we offset our response variable by fish standard length. We chose to offset our data by fish size as there is a known positive relationship between fish gape size and fish length, and gape is correlated with the size and type of prey a fish can consume ((Karpouzi \& Stergiou 2003, Barnes et al. 2010). We also accounted for site differences by including a random effect of site. We did not include an effect of season because English sole diet, which predominantly consists of polychaetes and bivalves, is consistent year-round (Reum \& Essington 2008). Because our response variable was a proportion, we implemented a betatrans function (Smithson \& Verkuilen 2005) on the response variable, and used a beta distribution in our model. The models were implemented using the glmmADMB() function in the R package glmmADMB (Fournier et al. 2012, Skaug et al. 2016) [Eq 4] for all prey items except brittle stars. The brittlestar model was implemented with the $\operatorname{glmmTMB}()$ function in the R package glmmTMB, year was scaled to our response variable and the model was fit with a beta family and link logit function. This function was used because the brittlestar dataset contained all zeros except for the year 2018. Because many statistical tests were conducted, we applied a False Discovery Correction to all p-values to reduce type I error (Benjamini \& Hochberg 1995). This correction was implemented using the padjust() function in base R.

Prey Proportional Mass $\sim$ Year Collected + fish standard length $+(1 \mid$ site $)$
[Eq. S1]

## Results

There was no change in the proportion of any prey item found in English sole (See Table S1 and Figure S1 Below).

Table S1. Results of the generalized linear mixed models for each prey item classified.

| Food Item | Predictor | estimate | std.error | statistic | p.value | p.bh |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Amphipod | Year | -0.003 | 0.005 | -0.552 | 0.581 | 1.000 |
| Amphipod | SL | 0.001 | 0.004 | 0.417 | 0.677 | 1.000 |
| Algae | Year | 0.002 | 0.005 | 0.530 | 0.596 | 1.000 |
| Algae | SL | 0.000 | 0.003 | -0.065 | 0.948 | 1.000 |
| Brachiopod | Year | 0.000 | 0.002 | 0.000 | 1.000 | 1.000 |
| Brachiopod | SL | 0.000 | 0.002 | 0.000 | 1.000 | 1.000 |
| Brittlestar | Year | 0.007 | 0.004 | 1.768 | 0.077 | 1.000 |
| Brittlestar | SL | -0.001 | 0.003 | -0.516 | 0.606 | 1.000 |
| Clam | Year | 0.007 | 0.005 | 1.504 | 0.133 | 1.000 |
| Clam | SL | 0.001 | 0.003 | 0.267 | 0.790 | 1.000 |
| Copepod | Year | 0.000 | 0.002 | 0.000 | 1.000 | 1.000 |
| Copepod | SL | 0.000 | 0.002 | 0.000 | 1.000 | 1.000 |
| Corophid Amphipod | Year | 0.000 | 0.002 | 0.000 | 1.000 | 1.000 |
| Corophid Amphipod | SL | 0.000 | 0.002 | 0.000 | 1.000 | 1.000 |


| Crab | Year | 0.003 | 0.004 | 0.682 | 0.495 | 1.000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Crab | SL | -0.001 | 0.003 | -0.296 | 0.767 | 1.000 |
| Crangon spp. | Year | -0.005 | 0.004 | -1.097 | 0.273 | 1.000 |
| Crangon spp. | SL | -0.002 | 0.003 | -0.507 | 0.612 | 1.000 |
| Cumacea | Year | -0.003 | 0.004 | -0.716 | 0.474 | 1.000 |
| Cumacea | SL | -0.002 | 0.003 | -0.757 | 0.449 | 1.000 |
| Isopod | Year | -0.002 | 0.003 | -0.569 | 0.570 | 1.000 |
| Isopod | SL | 0.000 | 0.002 | -0.107 | 0.914 | 1.000 |
| Mysid | Year | -0.003 | 0.004 | -0.813 | 0.416 | 1.000 |
| Mysid | SL | -0.001 | 0.003 | -0.347 | 0.728 | 1.000 |
| Polychaete | Year | 0.000 | 0.004 | -0.055 | 0.956 | 1.000 |
| Polychaete | SL | 0.004 | 0.003 | 1.151 | 0.250 | 1.000 |
| Scallop | Year | 0.002 | 0.005 | 0.530 | 0.596 | 1.000 |
| Scallop | SL | 0.000 | 0.003 | -0.065 | 0.948 | 1.000 |
| Snail | Year | 0.000 | 0.002 | 0.000 | 1.000 | 1.000 |
| Snail | SL | 0.000 | 0.002 | 0.000 | 1.000 | 1.000 |
| Tubeworm | Year | 0.001 | 0.005 | 0.308 | 0.758 | 1.000 |
| Tubeworm | SL | 0.003 | 0.003 | 0.810 | 0.418 | 1.000 |


|  |  |  |  |  |  |  |
| :--- | :--- | ---: | ---: | ---: | ---: | ---: |
| Worm | Year | -0.002 | 0.003 | -0.488 | 0.626 | 1.000 |
| Worm | SL | 0.002 | 0.003 | 0.708 | 0.479 | 1.000 |



Figure S1. The mean proportion of each type of English sole stomach content identified by decade.

| Table S2. List of UWFC specimens used in analyses. |  |  |
| :---: | :---: | :---: |
| Gut Content Analysis | CSIAA-N | Both |
| UW 048204 HSP 026 | UW 026114 | UW 10394_02 |
| UW 048204 HSP 033 | UW 026115_02 | UW 10394_03 |
| UW 048204 HSP 034 | UW 026118_01 | UW 110373_02 |
| UW 048204 HSP 036 | UW 026121 | UW 110373_05 |
| UW 048204 HSP 028 | UW 048602_01 | UW 110373_07 |
| UW 048204 HSP 029 | UW 10394_01 | UW 111168_01 |
| UW 18222_01 | UW 111197 | UW 112218 |
| UW 5815 | UW 112208_01 | UW 1177_01 |
| UW 4495_01 | UW 119958 | UW 151702_06 |
| UW 48351_01 | UW 119986_01 | UW 151702_07 |
| UW 48351_03 | UW 152939_01 | UW 151702_08 |
| UW 4885_01 | UW 155836_01 | UW 155822 |
| UW 5468_01 | UW 158625_01 | UW 16616_02 |
| UW 112224 | UW 158631_01 | UW 17077_01 |
| UW 026122 | UW 16925 | UW 17077_02 |
| UW 17077_02 | UW 5441_01 | UW 2269_02 |


| UW 110373_01 | UW 5523_01 | UW 40664_01 |
| :---: | :---: | :---: |
| UW 110373_03 | UW 5523_02 | UW 4491_01 |
| UW 110373_04 | UW 5829 | UW 4579_01 |
| UW 110373_06 | UW 789_01 | UW 48351_02 |
| UW 5441_01 |  | UW 48351_04 |
| UW 16925 |  | UW 4925_02 |
| UW 4614_01 |  | UW 6013_01 |
| UW 4794_01 |  |  |
| UW 4925_01 |  |  |
| UW 16589 |  |  |
| UW 2269_01 |  |  |
| UW 026120 |  |  |
| UW 110086_01 |  |  |
| UW 026114 |  |  |
| UW 026131_01 |  |  |
| UW 026131_02 |  |  |
| UW 026131_03 |  |  |
| UW 026115_03 |  |  |


| UW 026115_02 |  |  |
| :--- | :--- | :--- |
| UW 026115_01 |  |  |
| UW 026121 |  |  |
| UW 110080 |  |  |
| UW 16616_01 |  |  |
| UW 4885_02 |  |  |
| UW 10394_01 |  |  |
| UW 17851_01 |  |  |
| UW 17851_02 |  |  |
| UW 110312_01 |  |  |

## References

Barnes C, Maxwell D, Reuman DC, Jennings S (2010) Global patterns in predator-prey size relationships reveal size dependency of trophic transfer efficiency. Ecology 91:222-232 PubMed doi:10.1890/08-2061.1

Benjamini Y, Hochberg Y (1995) Controlling the False Discovery Rate: A Practical and Powerful Approach to Multiple Testing. J R Stat Soc Series B Stat Methodol 57:289-300 https://doi.org/10.1111/j.2517-6161.1995.tb02031.x

Fournier DA, Skaug HJ, Ancheta J, Ianelli J and others (2012) AD Model Builder: using automatic differentiation for statistical inference of highly parameterized complex nonlinear models. Optim Methods Softw 27:233-249 doi:10.1080/10556788.2011.597854

Karpouzi VS, Stergiou KI (2003) The relationships between mouth size and shape and body length for 18 species of marine fishes and their trophic implications. J Fish Biol 62:13531365 doi:10.1046/j.1095-8649.2003.00118.x

Reum JCP, Essington TE (2008) Seasonal Variation in Guild Structure of the Puget Sound Demersal Fish Community. Estuaries Coasts 31:790-801 doi:10.1007/s12237-008-9064-5

Skaug H, Fournier D, Nielsen A, Magnusson A, Bolker B (2016) Glmmadmb. R package version 0.8. 3.3

Smithson M, Verkuilen J (2005) Beta regression: practical issues in estimation. Disponıvel em http://www anu edu au/psychology/people/smithson/details/betareg/betareg html

## SUPPLEMENT 2

## Fish collection sites

Whenever possible, fish collection sites were reported as the latitude and longitude given from the UWFC record cards. If site names but no latitude and longitude were reported on a record card, we assigned a latitude and longitude coordinate to each unique site. If sites had similar names and were obviously identical in general locality, they were assigned the same latitude and longitude. We restricted the region in which we sampled to ensure our specimens were collected from Puget Sound proper and not the coast. Thus, we sampled specimens that were collected no further west than Port Angeles, WA and no further north than Bellingham, Bay, WA.

## Compound-specific stable isotope analysis of nitrogen amino acids (CSSIA-N)

Stepwise derivatization, analysis, and drift correction procedures are fully detailed in Welicky et al. (2021) and were based on Metges et al. (1996), Popp et al. (2007), and Chikaraishi et al. (2009). Briefly, we conducted CSSIA-N using a Trace 1310 GC in combination with a TriPlus RSH autosampler. All eluting compounds $\left(\mathrm{H}_{2} \mathrm{O}, \mathrm{CO}_{2}\right.$, and $\left.\mathrm{N}_{2}\right)$ off the column were oxidized inside a GC Isolink II combustion interface. Water was removed through a Nafion membrane downstream of the reactor while $\mathrm{CO}_{2}$ was cryogenically trapped in tubing submerged in liquid nitrogen before transfer to the IRMS (DELTA V) through a Conflo IV universal interface. All equipment was manufactured by Thermo Scientific, USA. High purity $\mathrm{N}_{2}$ ( $>99.9997 \% \mathrm{~N}_{2}$, Airgas) was used as reference gas to initially calculate the isotopic composition. Raw data were drift-corrected. See Welicky et al. (2021) for drift corrections procedures. Partially-automated drift correction procedures in R are available in a GitHub repository (https://github.com/rlwelicky/CSIA Allfish).

## Statistical Analysis

Testing for spatial autocorrelation Because some of the sites from which fish had been collected were close in space (geographically), we examined whether our data were spatially autocorrelated. To test for spatial autocorrelation, we first conducted generalized linear models with a Gaussian distribution, where the dependent variable was trophic position, glutamic acid, or phenylalanine and the independent variable was collection year. This was implemented using the 'glm' function in the "MASS" package (Venables \& Ripley 2020). Residuals of these models were extracted and then tested for spatial autocorrelation using Moran's I test for spatial autocorrelation using the function "testSpatialAutocorrelation" via the packages "DHARMa" (Hartig) and "spdep" (Bivand et al. 2013). We detected no spatial autocorrelation.

Testing for temporal autocorrelation Since our response variables may be influenced by factors that change in relation to time (i.e. season or annual patterns), we also examined whether our
data were temporally autocorrelated. We conducted a Durbin-Watson test to test for temporal autocorrelation for each of our response variables: trophic position, glutamic acid, and phenylalanine. We specified the order as a collection year and implemented this test using the 'dwtest' function from the package 'lmtest' (Zeileis \& Hortorn 2002). We detected no temporal autocorrelation. We could not assess seasonal autocorrelation as we did not have data from multiple seasons or even the transition of one season in a single year in our dataset. Nonetheless, we were curious if we could capture any seasonal autocorrelation. We could not test this continuously, because this would incorrectly sequence seasons, as 1 is January, and furthest numerically is 12 , December, yet seasonally these months and numbers are most similar. Therefore, we conducted a simple categorical regression analysis to test whether year was a significant predictor of season, and we determined there was no significant effect of year on month ( $\mathrm{est}=0.009, \mathrm{se}=0.007, \mathrm{t}=1.289, \mathrm{p}=0.199$ ). Testing for collinearity between fish size and year
Verifying if there was correlation between year and fish size This was implemented using the 'cor.test' function in base R, where the response variable was fish standard length and the independent variable was year. We determined that for Pacific Herring there was significant correlation between year and standard length, and there was no temporal trend in size for the other species sampled. Therefore after conducting our herring models, we verified that the variance inflation factors of the models were under 5 . We tested for collinearity using the check_collinearity function in the R package performance (Hebbali 2020). Acceptable value thresholds for the variance inflation factor (VIF) have been suggested to be 5 and 10 (reviewed in Dormann et al. 2013) and the VIF value for our models were 1.2 and under. Therefore, we did not detect collinearity in our models and no statistical corrections for collinearity were made. In a post hoc analysis, we verified that there was no significant interaction effect between size and year on rockfish trophic position ( $\mathrm{p}=0.276$ ) using the following model: Trophic position ${ }_{i j k} \sim$ Year collected ${ }_{j}+$ fish standard length ${ }_{i j k}+\left(1 \mid\right.$ site $\left._{k}\right)+$ Year collected ${ }_{j} *$ fish standard length ${ }_{i j k}$, where the response variable $e_{i j k}$ represents the trophic position value of the $i$ th fish collected from the $k$ th site in the $j$ th year. Year collected and fish standard length were centered and scaled.

Testing for discrete time points when rapid trophic change occurred We chose the recursive residual options because this option is better at detecting regime shifts than ordinary least squares residuals (e.g. Brown et al. 1975, Galpin \& Hawkins 1984). We used cumulative sums, rather than moving sums, because cumulative sums have been used in a number of disciplines to track and monitor total changes over time (e.g., Zeileis 2005, 2006). Cumulative sums also have been reported to perform better than moving sums (e.g., Brown et al. 2002), because moving sums can be sensitive to stochastic events (e.g., Zeileis \& Hortorn 2002).

## Supplemental Figures



Figure S2. Glutamic acid values across time by fish species. A = Pacific herring, B = English sole, $\mathrm{C}=$ walleye pollock and $\mathrm{D}=$ Pacific hake.


Figure S3. Phenylalanine values across time by fish species. $\mathrm{A}=$ Pacific herring, $\mathrm{B}=$ English sole, $\mathrm{C}=$ walleye pollock and $\mathrm{D}=$ Pacific hake.

## References

Bivand RS, Pebesma E, Gomez-Rubio V (2013) Applied spatial data analysis with R, Second edition. Springer, NY. http://www.asdar-book.org/
Brown RL, Durbin J, Evans JM (1975) Techniques for Testing the Constancy of Regression Relationships Over Time. J R Stat Soc Series B Stat Methodol 37:149-163 https://doi.org/10.1111/j.2517-6161.1975.tb01532.x
Brown SM, Benneyan JC, Theobald DA, Sands K and others (2002) Binary cumulative sums and moving averages in nosocomial infection cluster detection. Emerg Infect Dis 8:14261432 doi:10.3201/eid0812.010514

Chikaraishi Y, Ogawa N O, Kashiyama Y, Takano Y and others (2009) Determination of aquatic food-web structure based on compound-specific nitrogen isotopic composition of amino acids. Limnol Oceanogr Methods 7:740-750 doi.org/10.4319/lom.2009.7.740
Dormann CF, Elith J, Bacher S, Buchmann C and others (2013) Collinearity: a review of methods to deal with it and a simulation study evaluating their performance. Ecography 36:27-46 doi:10.1111/j.1600-0587.2012.07348.x

Galpin JS, Hawkins DM (1984) The Use of Recursive Residuals in Checking Model Fit in Linear Regression. Am Stat 38:94-105

Hebbali A (2020) Olsrr: Tools for Building OLS Regression Models.
Metges CC, Petzke KJ, Hennig U (1996) Gas chromatography/combustion/isotope ratio mass spectrometric comparison of N -acetyl-and N-pivaloyl amino acid esters to measure 15 N isotopic abundances in physiological samples: a pilot study on amino acid synthesis in the upper gastro-intestinal tract of minipigs. J Mass Spectrom 31:367-376 doi:10.1002/(SICI)1096-9888(199604)31:4<367::AID-JMS310>3.0.CO;2-V

Popp BN, Graham BS, Olson RJ, Hannides CCS and others (2007) Insight into the Trophic Ecology of Yellowfin Tuna, Thunnus albacares, from Compound-Specific Nitrogen Isotope Analysis of Proteinaceous Amino Acids. In: Terrestrial Ecology. Elsevier, p 173-190.

Venables WN, Ripley BD (2020) Modern Applied Statistics with S.
Welicky RL, Rolfe T, Leazer K, Maslenikov KP, Tornabene L, Holtgrieve GW, Wood CL (2021) Fluid-preserved fishes are one solution for assessing historical change in fish trophic level. Ecol Evol 11:415-426 doi:10.1002/ece3.7061

Zeileis A (2005) A Unified Approach to Structural Change Tests Based on ML Scores, F Statistics, and OLS Residuals. Econom Rev 24:445-466 doi:10.1080/07474930500406053

Zeileis A (2006) Implementing a class of structural change tests: An econometric computing approach. Comput Stat Data Anal 50:2987-3008 doi:10.1016/j.csda.2005.07.001
Zeileis A, Hortorn T (2002) Diagnostic Checking in Regression Relationships. R News 2:7-10

